

Homework 5. Due March 25

- (5 pts)** The Metropolis-Hastings algorithm is a modification of the Metropolis algorithm in which the transition matrix Q generating proposed moves does not need to be symmetric (see the last paragraph in Section 3.1 in `markov_chains.pdf`). Prove that the frequencies of visits of states in the Metropolis-Hastings algorithm approach the invariant distribution.
- (5 pts)** A random variable $\eta : \Omega \rightarrow [0, \infty]$ has an exponential distribution if

$$\mathbb{P}(\eta > t) = e^{-\lambda t} \quad \text{for all } t \geq 0,$$

where $\lambda \geq 0$ is a parameter.

Prove that a random variable $\eta : \Omega \rightarrow (0, \infty]$ has an exponential distribution if and only if it has the following memoryless property:

$$\mathbb{P}(\eta > t + s \mid \eta > s) = \mathbb{P}(\eta > t) \quad \text{for all } s, t \geq 0. \quad (1)$$

Hint: To deduce an exponential distribution from Eq. (1), you may proceed as follows. Introduce

$$F(t) := \mathbb{P}(\eta > t) \quad \text{and define } \lambda := -\log F(1).$$

Then write $F(1) = F(\frac{1}{n} + \dots + \frac{1}{n})$ and apply the memoryless property. Establish the exponential form for F for all rational t . Then use the fact that any real number can be approximated by rational numbers. Etc.

- (5 pts)** Consider a continuous-time Markov chain with a generator matrix L with all diagonal entries being equal to $-\lambda$ where $\lambda > 0$. Prove that the number of jumps occurring in the corresponding Markov jump process in a time interval $[0, t]$ is distributed according to the Poisson distribution with parameter λt , i.e.,

$$\mathbb{P}(J_k \leq t, J_{k+1} > t) = e^{-\lambda t} \frac{(\lambda t)^k}{k!}.$$

Hint: First show that for a small time interval $[0, h]$, $\mathbb{P}(J_1 \leq h, J_2 > h) = \lambda h + o(h)$ and $\mathbb{P}(J_2 \leq h) = o(h)$ where $o(h)$ denotes any function $f(h)$ such that $\lim_{h \rightarrow 0} h^{-1}f(h) = 0$. Then partition the interval $[0, t]$ into n subintervals where n is large. Show that $\mathbb{P}(J_k \leq t, J_{k+1} > t) = \mathbb{P}(A) + \mathbb{P}(B)$ where A is the event that exactly k subintervals contain one jump and the rest contain no jumps, and B is the event that at least one of the subintervals contains more than one jump and there are k jumps in total. Let n tend to ∞ and prove that $\mathbb{P}(B) \rightarrow 0$ while $\mathbb{P}(A) \rightarrow e^{-\lambda t} (\lambda t)^k (k!)^{-1}$.